Markov Networks

- MNs belong to the class of probabilistic graphical models Undirected, acyclic graphs of random variables (1)
- Example: Random variables: X_i , $i = 1, \dots, 5$
- Cliques (usually, maximal) in the graph have associated potential functions Non-negative real functions



- No conditional probabilities, but initially, local, joint marginal potentials
- Here, three maximal cliques, and three maximal sectors.
- Combination of potentials define define Joint probability distribution

Cliques' potentials become factors of the global joint distribution

• Here, three potentials for three cliques:

 $\psi_1(x_1, x_2, x_3), \ \psi_2(x_2, x_3, x_4), \ \psi_3(x_3, x_5)$

 Potentials may have parameters, possibly unknown, so as probability distributions They could be learned from data



• Joint probability distribution (density) for variables in MN:

 $P(x_1,...,x_5) := \frac{1}{Z} \times \psi_1(x_1,x_2,x_3) \times \psi_2(x_2,x_3,x_4) \times \psi_3(x_3,x_5)$

• Z is the "partition function", a normalization factor to obtain a probability distribution

It has to be: $\sum_{x_1,\ldots,x_5} P(x_1,\ldots,x_5) = 1$

- Then: $Z := \sum_{x_1,...,x_5} \psi_1(x_1, x_2, x_3) \times \psi_2(x_2, x_3, x_4) \times \psi_3(x_3, x_5)$
 - Z for "Zustandsumme" in German: "sum over states" (roots in Statistical Mechanics, initially largely developed by German speaking scientists)

- RVs X_i take values on their domains Dom(X_i)
 They reflect outcomes from a real-valued random experiment
 They are defined on the sample space Ω in common
- Example: (cont.) Assume Bernoulli RVs: $Dom(X_i) = \{0, 1\}$

Random propositional features

• Potentials:



- 1. $\psi_1(x_1, x_2, x_3) :=$ total number of 1s taken by the variables E.g. $\psi_1(1, 0, 1) = 2$
- 2. $\psi_2(x_2, x_3, x_4) := x_2 + x_3 + x_4$ E.g. $\psi_2(1, 0, 0) = 1$
- 3. $\psi_3(x_3, x_5) := x_3 \times x_5$

E.g. $\psi_3(0,1) = 0$

• Exercise: Compute Z above, the density value P(1, 0, 0, 1, 1), and the marginal value $P_{X_1}(1)$

For Z, compute the terms of the summation: (2^5 products)

1.
$$\psi_1(0,0,0) \times \psi_2(0,0,0) \times \psi_3(0,0) = 0 \times 0 \times 0 = 0$$

- 2. $\psi_1(0,1,1) \times \psi_2(1,1,1) \times \psi_3(1,1) = 2 \times 3 \times 1 = 6$
- 3. $\psi_1(0,1,0) \times \psi_2(1,0,0) \times \psi_3(0,0) = 1 \times 1 \times 0 = 0$, etc.

 $P(1,0,1,0,1) := \frac{1}{Z} \times \psi_1(1,0,1) \times \psi_2(0,1,0) \times \psi_3(1,1) = \frac{2 \times 1 \times 1}{Z}$

• There could be unknown parameters, to be learned, e.g.

 $\psi_2'(x_2, x_3, x_4) := lpha imes x_2 + (1 - lpha) imes x_3 + heta imes x_4$

- We have heard about the "Markov Condition", Markov Processes, etc.
- Main idea and intuition behind MNs:

The probability distribution of a particular variable (possibly with others in the net) depends only on a "small neighborhood" of the variable

There are implicit independence assumptions in place that "isolate" it from a large portion of the net

- The way MNs are constructed, via factorized representations, allows to identify certain stochastic (in)dependencies
- There are criteria to identify and exploit them

(notion of "d-separation")

Criteria also applicable to BNs

• A common class of MNs comes from Statistical Mechanics (SM): Boltzmann-Gibbs Distribution

 $P(\bar{x}) := \frac{1}{Z} \times exp(-\sum_{C} E(\bar{x}_{C})) = \frac{1}{Z} \times \Pi_{\bar{x}_{C}} \frac{1}{e^{E(\bar{x}_{C})}}$

Here, $\ \bar{x}$ represents the variables in the graph, and the \bar{x}_{C} those in clique C

A joint probability distribution from potentials: $\psi_{c} := \frac{1}{e^{E(\tilde{r}_{c})}}$

• Think of $E(\bar{x}_c)$ as an energy function of the variables of sub-state \bar{x}_c

This distribution makes low energy configurations (states) more likely

It penalizes high energy states

It favors higher entropy states (we will come back to this)

Energy function *E* may come in different forms
 Energy-based models are common in SM, Biochemistry, ML
 Whole families of distributions depending on the classes to which potentials belong

- MNs may be easier or more natural to use in some applications than BNs
- Choosing a direction between two variables may not be reasonable

E.g. in image analysis, with variables representing pixels of a same image

Also with relational data (think of attributes in a table)

- MNs have symmetries that BNs do not have, and can be exploited
- Inference with MNs tends to be more complex than with BNs

Some More Inference

- Let us see in more general terms what we did on page 12
- Idea: exploit distributive law $a \times b + a \times c = a \times (b + c)$ Three operations versus two
- Example: A chain model: $X_1 X_2 \cdots X_{N-1} X_N$ With potentials: $\psi(x_i, x_{i+1})$ Joint distribution: $P(\bar{x}) = \frac{1}{Z} \prod_{i=1}^{N-1} \psi(x_i, x_{i+1})$ Marginal of X_1 : $P_{X_1}(x_1) = \frac{1}{Z} \sum_{x_2, \dots, x_N} \prod_{i=1}^{N-1} \psi(x_i, x_{i+1})$
- Computed naively like this, the computation cost is proportional to Π^N_{i=1}|Dom(X_i)|
- By distributivity:

 $P_{X_1}(x_1) = \frac{1}{Z} \sum_{x_2} [\psi(x_1, x_2) \sum_{x_3} \psi(x_2, x_3) \cdots \sum_{x_{N-1}} \psi(x_{N-2}, X_{N-1}) \sum_{x_N} \psi(x_{N-1}, x_N)]]]$ Now cost proportional to $\sum_{i=1}^{N-1} |Dom(X_i)| \times |Dom(X_{i+1})|$ • Exercise: Consider the MN

Verify that:



 $P_{A}(a) := \frac{1}{Z} \sum_{b,c,d,e,f} \psi(a,b) \psi(a,c) \psi(b,d) \psi(c,e) \psi(b,e,f)$ $= \frac{1}{Z} \sum_{b} \psi(a,b) \sum_{c} \psi(a,c) \sum_{d} \psi(b,d) \sum_{e} \psi(c,e) \sum_{f} \psi(b,e,f)$

• This *variable elimination algorithm* uses distributivity Good for marginal of one variable • Example:

 $P(x_2) = \frac{1}{7} \sum_{x_1} \sum_{x_2} \sum_{x_3} \sum_{x_4} \psi(x_1, x_3, x_5) \psi(x_1, x_2) \psi(x_2, x_4) \psi(x_3, x_4)$ $O(2^5)$ operations in the naive way Х, X₂ with binary variables X₅ X4 X2 However: $P(x_2) = \frac{1}{Z} \sum_{x_1} \psi(x_1, x_2) \sum_{x_4} \psi(x_2, x_4) \sum_{x_3} \psi(x_3, x_4) \sum_{x_5} \psi(x_1, x_3, x_5)$ m₅ $= \frac{1}{Z} \sum_{x_1} \psi(x_1, x_2) \sum_{x_4} \psi(x_2, x_4) \sum_{x_3} \psi(x_3, x_4) m_5(x_1, x_3)$ m₂ $= \frac{1}{Z} \sum_{n} \psi(x_1, x_2) \sum_{x_1} \psi(x_2, x_4) m_3(x_1, x_4) \quad (m_i \text{ are marginals per clique or joins thereof})$ $= \frac{1}{Z}\sum_{x_1}\psi(x_1,x_2)m_4(x_1,x_2) = \frac{1}{Z}m_1(x_2)$ $O(2^3)$ now

Summing over x_2 gives Z (LHS is 1) ("messages" m_i could be reused, c.f. below) Not more that 3 variables appear together in any term of a summation

- In general, the maximum number of variables that appear together in a summation term depends on the elimination order
- The lowest complexity is obtained by the order that minimizes this maximum number

It is related to the tree-width of the graph



• Unfortunately, finding the optimal elimination order is NP-hard

Reduction from SAT

What about more than one marginal?

If we want more marginal distributions, we will be repeating operations

- The algorithm above can be adapted via reuse of precomputations
- There is a lot more about inference in PGMs ...

Tree-Width of a Graph

- Ţ
- The *tree-width* (TW) of a graph becomes relevant in many problems of data management and AI
- The TW of a graph measures how close a graph is to a tree
- It is commonly the case that graph problems become easier when the input graph has small TW
- Undirected graph $\mathcal{G} = \langle V, E \rangle$

A tree-decomposition of \mathcal{G} is a tree $\mathcal{T} = \langle \{S_1, \dots, S_n\}, E' \rangle$, such that:

- $S_1, \ldots, S_n \subseteq V$, i.e each node in \mathcal{T} is a subset of V
- $S_1 \cup \cdots \cup S_n = V$
- $(u, v) \in E \implies \{u, v\} \subseteq S_i$, for some *i*
- If for $v \in V$, $v \in S_j \cap S_k$, $i \neq k$, then $v \in S_i$, for every S_i in the unique (simple) path between S_j and S_k



- Width of tree decomposition \mathcal{T} : width $(\mathcal{T}) := (\max_i |S_i|) 1$
- The tree-width of graph *G*: tw(*G*) := min_τwidth(*T*)
 With *T* ranging over all tree decompositions of *G*
- When G is already a tree, the edges in E become the S_i



The S_i are connected by E' when they share a node in V





Chapter 6: Logical + Probabilistic KR

Leopoldo Bertossi

Probabilistic Approaches to KR

- Many logic-based approaches to KR&R have a probabilistic counterpart
- For example, a *default rule* (as in ASP) may be treated as a probabilistic/statistical statement
 As a conditional probability: *P(flies|bird)* = 0.95

"the probability of flying being a bird is 0.95"

- Consequences may be probabilistic too
- Diagnosis can be stated using conditionals: (by Bayes formula) $P(flu|fever) = \frac{P(flu) \times P(fever|flu)}{P(fever)}$ (a priori vs. a posteriori)
- More generally: $P(cause|symptom) = \frac{P(cause)P(symptom|cause)}{P(symptom)}$ P(symptom|cause) easier to estimate by experts than P(cause|symptom)

Probabilistic Reasoning Problems

- We can have PGMs or other probabilistic models
 With features that are random variables subject to some sort of uncertainty
- There are probabilistic approaches that favor representation of:
 - Joint distributions ~> "generative models"
 - MNs
 - Conditional distributions → "discriminative models"
 - BNs
 - Regression models: $Y = \alpha \times X + \beta + \epsilon$ Basically modeling P(Y|X)
- In principle, one can pass from one to the other, but there is complexity involved (remember inference)
 We did this with BNs, using the "chain rule" or Bayes formula

- Conditional probabilities allows us to attack several problems in *uncertain knowledge representation and reasoning*
- Probabilistic versions of diagnosis?

Consider an underlying probabilistic model \mathcal{K} (background knowledge) with an associated probability distribution $P_{\mathcal{K}}$

An observation O (or evidence), and a set of possible hypothesis (basic admissible explanations) $\mathcal{E} = \{E_1, \dots, E_n\}$

O is the value of a random variable (or several of them) in \mathcal{K} , and each E_i is (the value of) a random variable in \mathcal{K}

• We can attempt to find the *best explanation* $E^b \in \mathcal{E}$

$$E^{b} := \arg \max_{E \in \mathcal{E}} P_{\mathcal{K}}(E \mid O) \tag{1}$$

The most probable explanation given the evidence

• Usually called MAP-inference: *maximum a posteriori* After (conditioned on) the observation ... A different form of probabilistic reasoning: prefer an explanation E^{*}

$$E^{\star} := \arg \max_{E \in \mathcal{E}} P_{\mathcal{K}}(O \mid E)$$
 (2)

The explanation that maximizes the (conditional) probability of the observation

Which is what we observed after all ...

- This is similar to maximum-likelihood reasoning in Statistics
- <u>Exercise</u>: Verify that under the assumption that the explanations are equally likely (a priori), (1) reduces to (2) Hint: use Bayes formula
- There are model-dependent techniques for these reasoning tasks

Logic + Probability in Al

- Traditionally, the "logical-" and "probabilistic schools" have been separate and competitors
- In the last few years they have become complementary approaches
- Today, KR problems are attacked with mathematical models/techniques that involve simultaneously logic and probability
- Different forms of KR combine logic and probability for KR&R

Different formalisms, models, underlying assumptions, etc.

• These combined representations (models) can also be learned We will see some of them ...

- Conditional KBs: Knowledge base KB with
 - Hard knowledge, e.g. $emu \rightarrow bird$
 - Soft, conditional, probabilistic rules, of the form $r: (\alpha | \beta)[p]$ E.g. $r_{v}: (flies|bird)[0.9]$ (a "probabilistic conditional")
- Semantics? Logical consequences of/from KB?
- Possible-worlds semantics: Collection $\mathcal W$ of worlds $\mathcal W$
 - *W* is a set of propositional (or ground) atoms assumed to be true (Herbrand structures, as usual)
 - W must satisfy the hard knowledge in KB (as usual)
- W does not have to satisfy $\beta \to \alpha$, i.e. the conditional as a classical implication
- For this we need the probabilistic component ...

- We start considering a probability distribution P on W, the outcome space: $W \in W \mapsto P(W)$
- Which probability distribution *P*? (possibly several candidates)
- Since all the worlds in \mathcal{W} satisfy the hard knowledge, consider one that satisfies the conditionals:

For r: $(\alpha|\beta)[p]$, it must hold: $P(\alpha|\beta) = p$ (and $P(\beta) > 0$) meaning? $P(\alpha|\beta) := \frac{P(\alpha \land \beta)}{P(\beta)} := \frac{P(\{W \in W \mid W \models \alpha \land \beta\})}{P(\{W \in W \mid W \models \beta\})}$ (**)

Pick such a distribution P*

(which one?)

• Boolean query Q (expressed in the logical language): It may be true or false in an outcome world W

It becomes a Bernoulli RV: $P^{\star}(\mathcal{Q} (= 1)) := \sum_{\substack{W \in \mathcal{W} \\ W \models \mathcal{Q}}} P^{\star}(W)$

- Example: Propositional variables: yellow, fly, bird, emu, canary,...
 KB = {bird, emu → bird, (flies|bird)[0.9], canary → yellow, ...}
- $\mathcal W$ contains worlds satisfying the hard knowledge:

(logical constraints)

 $\begin{aligned} &W_1 = \{ \textit{yellow, bird, canary} \}, \\ &W_2 = \{ \textit{yellow, bird, fly, canary} \}, \\ &W_3 = \{ \textit{yellow, emu, bird, fly, canary} \}, \text{ etc.} \end{aligned}$

- Assume there is a distribution P on W
- Query Q: yellow \land bird \rightarrow fly? It is true in $W_2, W_3, ...$
- Event associated to the query: $E(Q) := \{W_2, W_3, ...\}$ $P(Q) := P(\{W_2, W_3, ...\}) = P(W_2) + P(W_3) + \cdots$

- More generally: We obtain formulas as consequences with associated probabilities
- We could also define the logico-probabilistic consequences of *KB* as those with high probability
- For a logical sentence φ (or query):

 $KB \models_{P} \varphi :\iff P(\varphi) > 1 - \epsilon$

As in the previous example, φ defines an event

- ϵ can be pre-specified (and small)
- Which is a good distribution P on W?

A preferred P^* ?

Some may be "better" or more justified than others

• <u>MEzzistributions</u>: Prefer a distribution that does not make unjustified, arbitrary assumptions

- One that does not impose unnecessary "structure or complexity" on the model
- Think of Statistical Mechanics: the contents of a gas container tends to reach a state of equilibrium of maximum disorder, with low complexity or structure
- The notion of Entropy comes in ...

Systems tend to reach equilibrium states of maximum entropy (maximum disorder)

To impose order, structure, complexity, one needs extra energy



(an unlikely state)

• Choose a distribution that maximizes the entropy?

- Entropy: Probability space $\langle \Omega, P \rangle$, with $\Omega = \{\omega_1, \dots, \omega_n\}$, $p_i := P(\omega_i)$ (finite case for simplicity)
- Entropy of the distribution:

$$Entropy(P) := -\sum_{i=1}^{n} p_i \times log(p_i) \qquad (*)$$
$$= \sum_{i=1}^{n} (p_i \times log(\frac{1}{p_i})) \qquad (= H(P))$$

• Entropy is interpreted as a measure of the level of uncertainty captured by the distribution

A measure of the degree of disorder it attributes to the system

- This "measure" can be derived from some desirable properties As the only function that satisfies them (a theorem)
- Furthermore, one can prove: The uniform distribution maximizes the entropy, i.e. p_i = ¹/_n
 When there is no extra constraint to satisfy or knowledge to consider

• Back to our problem, it makes sense to choose P^* as the maximum-entropy distribution:

> $P^{\star} := \arg \max_{P \in \mathcal{P}} Entropy(P)$ = arg max_{ber} $-\sum_{W \in \mathcal{W}} P(W) \times ln(P(W))$

• Conditioned maximization problem over the class ${\cal P}$ of probabilities that satisfy the conditions above (c.f. page 8)



- Distribution without arbitrary assumptions/structure, maximum disorder, maximum independence
- Choose a distribution that is as close to the uniform distribution as possible given the conditions The one that is the least unjustified ...
- One can define query answering and logico-probabilistic consequences from KB as on pages 8 and 12

- Example: Consider a box containing balls and cubes, which can be white or green. We know that all balls are white. Possible distributions?
- We can think this scenario as involving a draw from the box, whose observation gives rise to 2 random variables (features) *Shape, Color,* each taking two values
- Joint distribution *P*(*Shape*, *Color*) under conditional *P*(*Color* = *g*|*Shape* = *b*) = 0?

1.	Dist	w Bs	g Bs	w Cs	g Cs	Entropy (in bits)
	1	$\frac{1}{5}$	0	2 5	2 5	? (compute)

Assuming that 20% of objects are balls

This entails a lot about color, and shape given color: (check!) $w = \frac{3}{5}, g = \frac{2}{5}, b|w = \frac{1}{3}, c|w = \frac{2}{3}, b|g = 0, c|g = 1$ $P(g, b) = 0, P(w, b) = \frac{1}{5}, P(g, c) = \frac{2}{5}, P(w, c) = \frac{2}{5}$

\mathbf{c}	Dist	w Bs	g Bs	w Cs	g Cs	Entropy (in bits)
2.	2	$\frac{2}{5}$	0	2 5	$\frac{1}{5}$? (compute)

Assuming 20% of objects to be green, which leads to: (check!) $w = \frac{4}{5}, g = \frac{1}{5}, b|w = \frac{1}{2}, c|w = \frac{1}{2}, b|g = 0, c|g = 1$

3.	Dist	w Bs	g Bs	w Cs	g Cs	Entropy (in bits)
	3	$\frac{1}{3}$	0	$\frac{1}{3}$	$\frac{1}{3}$? (compute)

No assumption determining other properties

Dist	w Bs	g Bs	w Cs	g Cs	Entropy (in bits)
1	$\frac{1}{5}$	0	$\frac{2}{5}$	$\frac{2}{5}$	1.522
2	$\frac{2}{5}$	0	$\frac{2}{5}$	$\frac{1}{5}$	1.522
3	$\frac{1}{3}$	0	$\frac{1}{3}$	$\frac{1}{3}$	1.585

Last row corresponds to maximum entropy distribution ...

Markov Logic Networks

- MLNs combine FO logic and Markov Networks (MNs) in the same logico-probabilistic representation
- They are used for uncertain Knowledge Representation and Reasoning, and also in Machine Learning
 Networks can be learned from data for producing KR models, with new forms of inference
- MLNs belong to Statistical Relational Learning (SRL) Handling inherent uncertainty and exploiting compositional structure are fundamental to understanding and designing large-scale systems

Statistical relational learning builds on ideas from probability theory and statistics to address uncertainty while incorporating tools from logic, databases, and programming languages to represent structure

• We have a knowledge base *KB* in FO logic, but formulas have "weights" (eventually leading to probabilities)

- Ground atoms of the logical language become the nodes in an undirected graph that is handled as a MN
- The formulas can be used to define cliques, and their weights to define potentials on cliques, and so on ...

• Example: (a simplified form of MLN)

Consider the implicitly universally quantified constraint (w/variables)

3.9: $Manager(M, E) \rightarrow HighlyCompensated(M)$ (*)

• Consider all possible ground atoms built with underlying domain *Dom*

 $Atoms_{Dom} = \{Man(m, e) \mid (m, e) \in Dom \times Dom\} \cup \{HC(m) \mid m \in Dom\}$

- Each of these ground atoms becomes a node in a MN
- More precisely, each atom A ∈ Atoms_{Dom} becomes a Bernoulli random variable X_A in the MN (it can be true or false)
- These variables are stochastically and mutually dependent with (some of the) other variables X_{A'}
 This will be determined by the edges and potentials in a MN
- The MN has a set of nodes V of size $M = 4^2 + 4$ nodes

• The groundings of the MLN are: (4² of them)

1.	$ eg M(d_1, d_1) \lor HC(d_1)$		
2.	$ eg M(d_1, d_2) \lor HC(d_1)$	(<i>F</i> ₂)	$\overline{\mathbf{r}}$

16. $\neg M(d_4, d_4) \lor HC(d_4)$

- Each grounding represents a factor in the underlying MN:
- The instantiations 1.-16. of (*) become the factors

E.g. the factor or clique F_2 : $M(d_1, d_2) - HC(d_1)$ in the MN

This will be a (mini) clique which will have an associated potential depending on its weight

- Weight $w(F_2) = 3.9$ (inherited from weight for original formula)
- We do not have potentials yet, only the graph
- Weight of a factor determines potential of associated clique

$$\psi_{F_2}(\underbrace{\mathcal{M}(d_1, d_2)}_{x_1}, \underbrace{\mathcal{HC}(d_1)}_{x_2}) \xrightarrow{(x_1, x_2)}_{x_2} \stackrel{(x_1, x_2) :=}{\underset{3.9 \text{ otherwise}}{\text{ if } x_1 \xrightarrow{\forall x_1 \text{ and } x_2 \xrightarrow{\forall x_2 \text{ i.e. } F_2 \text{ false}}}_{\text{i.e. } F_2 \text{ false}$$

- Similarly for the other 15 factors (original weight inherited by factors)
- Product of potentials defines distribution P^m over possible worlds
 Indirectly over the Bernouilli RVs X_i (normalized product of their potentials)
- A possible world $W_1 = \{M(d_1, d_2), M(d_3, d_1), HC(d_1), HC(d_4)\}$
- W_1 makes true all factors, except for $\neg M(d_3, d_1) \lor HC(d_3)$
- In compatibility with MNs, its weight (or joint potential): $weight(W_1) := \prod_{F:W_1 \models F} w(F) = (3.9)^{15}$

Product of the weights of the factors that are true in W_1

- Probability of W_1 : $P^m(W_1) := \frac{weight(W_1)}{7}$ (also from (*))
- Normalization denominator: $Z = \sum_{worlds w} weight(W)$
- <u>Exercise</u>: How large is the number *M* of nodes in the MN depending on the size *n* of *Dom* and the predicates?
- Let see now a more common way of presenting MLNs

• Example: Real-valued weight $w(\varphi)$ assigned to formulas $\overline{\varphi \in KB}$

Formula	Weight
$\forall x(Steal(x) \rightarrow Prison(x))$	3
$\forall x \forall y (CrimePartners(x, y) \land Steal(x) \rightarrow Prison(y))$	1.5
	•••

- Fixed, finite domain, e.g. $Dom = \{bob, anna, ...\}$
- Producing ground atoms, e.g. CrimePartners(bob, anna), and instantiated formulae, e.g. Steal(bob) → Prison(bob)
- Edge between two nodes (ground atoms) if they appear in a same instantiated formula



A (local, mini) clique for one instantiation of the second formula $% \left(\left[\left({{{\left({{{\left({\left({{\left({{{\left({{{\left({{\left({{\left({{{\left({{{\left({{{\left({{{\left({{{\left({{{\left({{{\left({{{\left({{{}}}}} \right)}}}}\right,}$

 As on many occasions so far, a world is a set of ground atoms A Herbrand structure indicating what is true (and indirectly what is not)

 $W_1 = \{CrimePartners(bob, anna), Steal(bob)\}$

 $W_2 = \{CrimePartners(bob, anna), Steal(bob), Prison(anna)\}$

- A world may satisfy an instantiation of a formula or not For example, W₂ satisfies "the clique" above, but W₁ not
- The higher the weight, the higher the difference between a world that satisfies the formula and one that does not (with the rest the same)
- The worlds get associated probabilities through the weights
- A world that violates a formula is not invalid (not non-model), but only less likely

Some "models" (worlds) become more likely than others

- The weight of a formula captures the way the probability decreases when a ground instance of the formula is violated
- A high weight for a formula becomes a high penalty on worlds that do not satisfy it
- Given a world W, each node N ∈ V takes the value 0 or 1 if false or true in W (worlds become outcomes) Then, each node N becomes a Bernoulli random variable X^N
- Worlds become instantiations of a random vector

 $\mathcal{X} = \langle X^{N_1}, X^{N_2}, X^{N_3}, \dots, X^{N_M} \rangle$



 W_1 becomes $\mathbf{x}_1 = \langle 1, 1, 0, \dots, 0 \rangle$

- Each instantiation of a formula generates a propositional "feature", with value 1 if true in a world W, and 0, otherwise
- We can assign probabilities to worlds
 Equivalently, build a joint probability distribution P^m for X
- As with MNs, we can use a log-linear "potential function"
- For world W associated to $\mathbf{x} \in \{0, 1\}^M$: $P^m(W) := P^m(\mathcal{X} = \mathbf{x}) := \frac{1}{Z} \times e^{\sum_{\varphi \in KB} w(\varphi) \times n(\varphi, \mathbf{x})}$ (*)
- n(φ, x): number of instantiations of φ true in world x (or its clique x_C)
- Z normalizes over all possible worlds:

$$Z = \sum_{\mathbf{z} \in \{0,1\}^M} \exp(\sum_{\varphi \in KB} w(\varphi) \times n(\varphi, \mathbf{z}))$$

- From (*): A (ground) clique gc associated to a formula φ in the MN has the potential: ψ_{gc}(x) := exp(w(φ) × I_{gc}(x)), with x formed by 0s and 1s
- In the example, gc could be the three ground atoms in the top-left corner: gc = {N₁, N₂, N₃}

 $\exists g_{c}(\bar{x})$, the *indicator function*, takes value 1 if g_{c} true for \bar{x} , = and 0 otherwise (with that, $e^{0} = 1$ gives the right factor)

- This can be seen as a Gibbs distribution for MNs
- Since we divide by all possible satisfaction with possible worlds (the Z), we can see w(φ) as a penalty for not satisfying it Because in that case, it is multiplied by 0
- So, hard or strong constraints that we want to see satisfied should have high weights

- We obtain a probability distribution over possible worlds Those that satisfy "more" high-weight (instances of) formulas become more likely
- <u>Exercise</u>: Give an example of a MLN with a model that (logically) violates all the formulas *F* in KB, as universal ICs, but still has a non-zero probability

Hint: Make sure not all ground instantiations of the ICs become false

• With a MLN we do not have to create the actual, underlying, ground MN

We have a pattern to produce a concrete one if needed

• Having the exponential on page 24 allows us to deal with sums instead of products

- It is possible to extend MLNs with functions symbols Using Skolem functions could be used for formulas with existential quantifiers
- One can learn MLNs

Learn the weights and/or the formulas

The latter define the structure of the underlying and implicit $\ensuremath{\mathsf{MN}}$

• How to do inference with MLNs?

Inference in Markov Logic Networks

- Inference under MLNs is of a probabilistic nature
- Similarly, the MLN defines a probability distribution *P^m* over the possible worlds
- Basic inference task is computing the probability of a world, as on page 24

More interesting is a query in the language of the KB: For a sentence ψ :

 $P^{m}(\psi) := P^{m}(\{W \mid W \models \psi\}) := P^{m}(\mathcal{X} \text{ makes } \psi \text{ true})$

 $= P^m(\{\mathbf{x} \in \{0,1\}^M \mid \psi \text{ is true in } \mathbf{x}\}) = \sum_{\mathbf{x} \in \{0,1\}^M \mid \psi \cdots} P^m(\mathbf{x})$

 Computing the probabilities amounts, directly or not, to counting models (possibly with specific properties)
 Here, a form of weighted model counting
 A hard computational problem ...

- In general in SRL, we want to avoid as much as possible doing the grounding of formulas
 Followed by the explicit weighted model counting (bound to be computationally complex)
- Can we stay at a higher ("lifted") level?
- Different areas converge: model counting in logic (around SAT-related problems), graph theory, and data management



E.g. in the presence of constraints (here a referential constraint)



- Too many variables and groundings, many not related to each other
- SRL is precisely about doing things at the higher, relational or FO logical level

Representation and reasoning at a "lifted", more general level of granularity

- Can we do model counting without instantiation?
- Can we approximate model counting (and probabilities) without instantiation?
- Doing what is called "Lifted Inference" Lifted up to the FO representation Exploiting patterns, independence and symmetries