

Some of the useful rules to compute probabilities

1. $\Pr(\emptyset) = 0$

2. $A_1, A_2, A_3, \dots, A_n$ are pairwise disjoint events, then

$$\Pr(A_1 \cup A_2 \cup \dots \cup A_n) = \sum_{i=1}^n \Pr(A_i)$$

3. For any event A , $\Pr(A) \equiv 1 - \Pr(\bar{A})$

4. If A and B are events $\Pr(A \cup B) = \Pr(A) + \Pr(B) - \Pr(A \cap B)$

5. UNION BOUND

If $A_1, A_2, \dots, A_n$ are a sequence of events, then

$$\Pr(A_1 \cup A_2 \cup \dots \cup A_n) \leq \sum_{i=1}^n \Pr(A_i)$$

6. A and B are events and $A \subseteq B$ then $\Pr(A) \leq \Pr(B)$

7. $\Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)} \iff \Pr(A \cap B) = \Pr(A|B) \Pr(B)$

8. A and B are independent if $\Pr(A \cap B) = \Pr(A) \Pr(B)$

9. $A_1, A_2, \dots, A_n$ are mutually independent if

$$\Pr(A_1 \cap A_2 \cap \dots \cap A_n) = \Pr(A_1) \Pr(A_2) \dots \Pr(A_n)$$

Conditional Probability

Example: Consider that a manufacturer produces 100 items of which 20 are defective. Suppose we choose 2 items from the lot in the following two ways:

Method 1: We choose the first item, put it back in the lot and then choose second (with replacement)

Method 2: We choose the first item, do not put it back, then choose the second. (without replacement)

Define the events.

$A = \{\text{the first item is defective}\}$

$B = \{\text{the second item is defective}\}$

For Method 1 (with replacement)

$$\Pr(A) = \Pr(B) = \frac{20}{100} = \frac{1}{5}$$

For Method 2 (without replacement)

$$\Pr(A) = \frac{20}{100} = \frac{1}{5}$$

What about $\Pr(B) = ?$

That depends on what was the first item?

If first item was non-defective $\Pr(B) \rightarrow \frac{20}{99}$

If first item was defective $\Pr(B) \rightarrow \frac{19}{99}$

Define

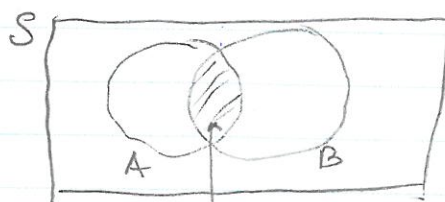
$\Pr(B|A) \equiv$ Conditional probability of the event B given that A has occurred.

Thus

$$\Pr(B|A) = \frac{19}{99} \rightarrow \begin{array}{l} \text{Only 19 defective items are left.} \\ \text{99 items are left after first item is removed} \end{array}$$

$$\Pr(B|\bar{A}) = \frac{20}{99}$$

Note: $\Pr(B|A)$ — We are computing $\Pr(B)$ with respect to a reduced sample space A , rather than with respect to full sample space S .



how probable are we in B , given that we are in A .

Example 2

Two fair dice are tossed.

Let x_1 be outcome of first.

Let x_2 be the outcome of second.

Define two events

$$A = \{(x_1, x_2) \mid x_1 + x_2 = 10\} \quad B = \{(x_1, x_2) \mid x_1 > x_2\}$$

Note that

$$\Pr(A) = \frac{3}{36} \quad \leftarrow \{(5,5), (4,6), (6,4)\}$$

$$\Pr(B) = \frac{15}{36} \quad \leftarrow \{(2,1), (3,1), (4,1), (5,1), (6,1), \dots\}$$

$$\Pr(A|B) = \frac{1}{15} \quad \leftarrow \{(6,4)\} \text{ from } \{(2,1), (3,1), \dots\}$$

$$\Pr(B|A) = \frac{1}{3} \quad \leftarrow \{(6,4)\} \text{ from } \{(5,5), (4,6), (6,4)\}$$

Let us look at $\Pr(A \cap B) = \frac{1}{36}$.

Now observe that

$$\Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)} = \frac{1/36}{15/36} = 1/15$$

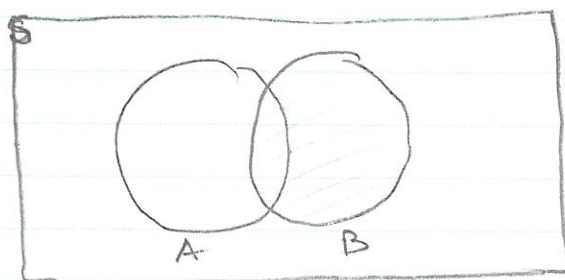
$$\Pr(B|A) = \frac{\Pr(A \cap B)}{\Pr(A)} = \frac{1/36}{3/36} = 1/3$$

Definition

Let $(S, \Pr)$ be a probability space, and let A and B be two events with $\Pr(B) > 0$.

The conditional probability $\Pr(A|B)$ is defined as

$$\Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)}$$



- Given B , event A occurs only if $A \cap B$ occurs.
- If both $\Pr(A) > 0$, $\Pr(B) > 0$, then: $\Pr(B|A) = \frac{\Pr(A \cap B)}{\Pr(A)}$.
- Note that from the definition we can

$$\text{derive that } \Pr(A \cap B) = \Pr(A|B) \cdot \Pr(B) \\ = \Pr(B|A) \cdot \Pr(A)$$

For example, consider Method 2, and we ask what is the probability that both items are defective, i.e. what is $\Pr(A \cap B) = ?$

$$\Pr(A \cap B) = \Pr(A|B) \cdot \Pr(B) = \Pr(B|A) \cdot \Pr(A)$$

$$= \frac{19}{99} \cdot \frac{20}{100}$$

$$= 19/495$$

() Claim: Let A and B be two events with $\Pr(B) > 0$.

$$\text{Then } \Pr(A|B) + \Pr(\bar{A}|B) = 1$$

$$\text{Proof: } \Pr(A|B) + \Pr(\bar{A}|B)$$

$$= \frac{\Pr(A \cap B)}{\Pr(B)} + \frac{\Pr(\bar{A} \cap B)}{\Pr(B)} \quad \text{--- (1)}$$

~~the~~ Events $A \cap B$ and $\bar{A} \cap B$ are disjoint

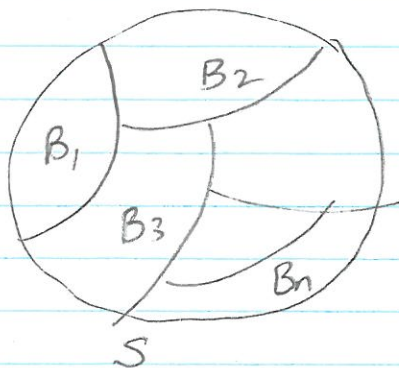
$$\text{thus } \Pr(A \cap B) + \Pr(\bar{A} \cap B) = \Pr(A \cap B \cup \bar{A} \cap B)$$

$$= \Pr(B)$$

Thus (1) ~~is~~ is same as

$$= \frac{\Pr(B)}{\Pr(B)} = 1. \quad \blacksquare$$

Next up: Law of Total Probability.



Assume that $B_1, B_2, \dots, B_n$ is a partition of S

i.e

$$(a) B_1 \cup B_2 \cup \dots \cup B_n = S$$

$$(b) \forall i, j, i \neq j, B_i \cap B_j = \emptyset.$$

Theorem [Law of total probability]

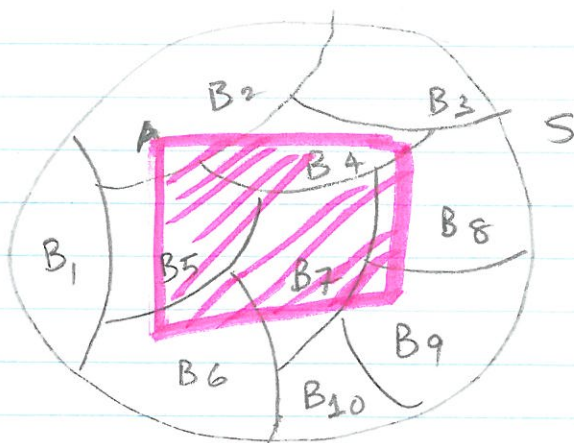
Let $(S, \Pr)$ be a probability space and let A be an event.

Assume that $B_1, B_2, \dots, B_n$ be events such that

- 1) $\Pr(B_i) > 0 \quad \forall i, \quad 1 \leq i \leq n$
- 2) the events $B_1, B_2, \dots, B_n$ are pairwise disjoint.
- 3) $\bigcup_{i=1}^n B_i = S$

Then

$$\Pr(A) \equiv \sum_{i=1}^n \Pr(A|B_i) \cdot \Pr(B_i)$$



Proof: $A = A \cap S$ as $A \subseteq S$.

$$= A \cap \left(\bigcup_{i=1}^n B_i \right)$$

$$= \bigcup_{i=1}^n (A \cap B_i)$$

Since $B_1, B_2, \dots, B_n$ are pairwise disjoint, then $A \cap B_1, A \cap B_2, \dots, A \cap B_n$ are pairwise disjoint.

Thus,

$$\Pr(A) = \Pr\left(\bigcup_{i=1}^n A \cap B_i\right)$$

$$= \sum_{i=1}^n \Pr(A \cap B_i)$$

$$= \sum_{i=1}^n \Pr(A|B_i) \Pr(B_i)$$

$$\text{as } \Pr(A|B_i) = \frac{\Pr(A \cap B_i)}{\Pr(B_i)}$$

by definition of conditional probability

Applications

Example 1:

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Experiment:

Flip a coin —
 Heads → Roll a ^{fair} die and let R be result
 Tails → Roll two ^{fair} dice. Let R be the sum.

Define the event

$$A = \{\text{the value of } R = 2\}$$

Problem: Compute $\Pr(A) = ?$

Define the event $B = \{\text{coin comes up heads}\}$.

Note ① $\Pr(B) > 0$, $\Pr(\bar{B}) > 0$

$$\text{② } B \cap \bar{B} = \emptyset$$

$$\text{③ } S = B \cup \bar{B} = \text{possible outcomes of toss}$$

Then,

$$\Pr(A) = \Pr(A|B) \Pr(B) + \Pr(A|\bar{B}) \Pr(\bar{B})$$

$$\Pr(A|B) = \begin{array}{l} \text{given that} \\ \text{Toss is head then what is} \\ \text{the probability we get a 2.} \end{array}$$

$$= 1/6$$

$$\Pr(A|\bar{B}) = 1/36$$

$$\text{Thus } \Pr(A) = 1/6 \cdot 1/2 + 1/36 \cdot 1/2 \equiv 7/72 \quad \blacksquare$$

Applications

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Example 2: Factory producing 100 items,

20 Defective, 80 non-defective. We choose two items without replacement.

Events $I = \{\text{first chosen item is defective}\}$

$\bar{I} = \{\text{Second chosen item is defective}\}$

Suppose we want to compute $\Pr(\bar{I}) = ?$

Let $B_1 = I$ and $B_2 = \bar{I}$

Note that $S = I \cup \bar{I}$ and $I \cap \bar{I} = \emptyset$
and $\Pr(I) > 0$, $\Pr(\bar{I}) > 0$

Let $B_1 = I$, $B_2 = \bar{I}$, and $A = \bar{I}$. and apply theorem

$$\Pr(\bar{I}) = \Pr(A) = \sum_{i=1}^2 \Pr(A|B_i) \cdot \Pr(B_i)$$

$$= \Pr(A|B_1) \cdot \Pr(B_1) + \Pr(A|B_2) \cdot \Pr(B_2)$$

$$= \frac{19}{99} \cdot \frac{1}{5} + \frac{20}{99} \cdot \frac{4}{5}$$

$$= \frac{1}{5}$$

(Probability that the second item is defective is $1/5$ under without replacement).

Example 3: An item is produced by 3-factories

F_1 produces $2x$ items

F_2, F_3 produce x items each.

	F_1	F_2	F_3
# items	$2x$	x	x
% defective	2%	2%	4%

All items are put in a pile and one is drawn at random.

Question: $\Pr(\text{drawn item is defective}) = ?$

$A = \{ \text{drawn item is defective} \}$

$B_1 = \{ \text{drawn item is from } F_1 \}$

$B_2 = \{ \text{drawn item is from } F_2 \}$

$B_3 = \{ \text{drawn item is from } F_3 \}$

Note that (1) $\Pr(B_1) > 0, \Pr(B_2) > 0, \Pr(B_3) > 0$

(2) $B_i \cap B_j = \emptyset \quad i \neq j \quad i, j \in \{1, 2, 3\}$

(3) $B_1 \cup B_2 \cup B_3 = S$ are all possibilities for the drawn item.

Thus

$$\Pr(A) = \Pr(A|B_1) \Pr(B_1) + \Pr(A|B_2) \Pr(B_2) + \Pr(A|B_3) \Pr(B_3)$$

$$= 0.02 * \frac{1}{2} + 0.02 * \frac{1}{4} + 0.04 * \frac{1}{4}$$

$$= 0.025 \quad \blacksquare$$

Question: Suppose that an item is chosen from the pile, and is found to be defective.

What is the probability that it is produced by F_1 .

$$\Pr(B_1|A) = ?$$

BAYES THEOREM

Let $B_1, B_2, \dots, B_n$ be partition of sample space S .
Let $A \subseteq S$ be an event.

$$\Pr(B_i|A) = \frac{\Pr(A|B_i) \Pr(B_i)}{\sum_{j=1}^n \Pr(A|B_j) \Pr(B_j)}$$

$$\begin{aligned} \Pr(B_1|A) &= \frac{\Pr(A|B_1) \cdot \Pr(B_1)}{\Pr(A|B_1) \Pr(B_1) + \Pr(A|B_2) \cdot \Pr(B_2) + \Pr(A|B_3) \cdot \Pr(B_3)} \\ &= \frac{0.02 * 1/2}{0.02 * 1/2 + 0.02 * 1/4 + 0.04 * 1/4} \\ &= 0.40 \end{aligned}$$

Note: Bayes thm allows us to write conditional probabilities in terms of similar conditional probabilities but the conditions are reversed.

Two events A and B are said

to be independent if $Pr(A \cap B) = Pr(A) Pr(B)$.

Example: Roll a fair die twice.

$A = \{ \text{the first roll shows an even number} \}$

$B = \{ \text{2nd Roll shows 5 or 6} \}$

$$Pr(A) = \frac{1}{2} \quad Pr(B) = \frac{1}{3} \quad Pr(A \cap B) = ? = \frac{1}{6} \\ = Pr(A) Pr(B)$$

Example 2: Roll two fair die

$A = \{ \text{the first die shows even number} \}$

$B = \{ \text{the second die shows an odd number} \}$

$C = \{ \text{two die shows both odd or both even numbers} \}$

Observe $Pr(A) = Pr(B) = Pr(C) = \frac{1}{2}$

$$Pr(A \cap B) = \frac{1}{4} = Pr(A) Pr(B)$$

$$Pr(A \cap C) = \frac{1}{4} = Pr(A) Pr(C)$$

$$Pr(B \cap C) = \frac{1}{4} = Pr(B) Pr(C)$$

} pairwise independent

$$Pr(A \cap B \cap C) = 0 \neq Pr(A) Pr(B) Pr(C)$$

~~Pair~~

Definition: 3 events A, B, C are mutually

independent iff

$$Pr(A \cap B) = Pr(A) Pr(B)$$

$$Pr(B \cap C) = Pr(B) Pr(C)$$

$$Pr(A \cap C) = Pr(A) Pr(C)$$

$$Pr(A \cap B \cap C)$$

$$= Pr(A) Pr(B) Pr(C)$$

Mutually Independent Events

Let (S, Pr) be a probability space.

A sequence of events $A_1, A_2, \dots, A_n$; $n \geq 2$;

is said to be Mutually Independent Events if

for all k , $2 \leq k \leq n$ and all

$$i_1 < i_2 < \dots < i_k$$

$$Pr(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_k}) = Pr(A_{i_1}) \cdot Pr(A_{i_2}) \cdot \dots \cdot Pr(A_{i_k})$$

As noted in previous example,

pairwise independence $\not\Rightarrow$ mutual independence.

Examples of mutual independence.

Question: Let $n \geq 1$ and we flip n -fair coins.

Define event $A_i = \{i^{\text{th}} \text{ coin comes up heads}\}$
for $1 \leq i \leq n$.

Define $A = A_1 \cap A_2 \cap A_3 \dots \cap A_n$, i.e. all n -flips are heads.

Clearly $\Pr(A) = 1/2^n$.

Alternatively, using independence, as each coin flip is independent of any other flip.

$$\Pr(A) = \Pr(A_1) \cdot \Pr(A_2) \cdot \dots \cdot \Pr(A_n) = \left(\frac{1}{2}\right)^n = \frac{1}{2^n}.$$

Question:

Consider a circuit C that consists of n -components $C_1, C_2, \dots, C_n$. Let p be a real number $0 \leq p \leq 1$.

Assume that any component may fail

with probability p , and its failure is indep of other components

Define events $1 \leq i \leq n$, where

$$A_i = \{ \text{"Component } C_i \text{ fails"} \}$$

Suppose C is configured as



C fails if any of C_i fails.

$$A = \{ \text{"the circuit } C \text{ fails"} \}$$

$$\Pr(A) = ?$$

Note that

$$A = A_1 \vee A_2 \vee A_3 \dots \vee A_n.$$

$$\Leftrightarrow \bar{A} = \bar{A}_1 \wedge \bar{A}_2 \wedge \bar{A}_3 \dots \wedge \bar{A}_n$$

Thus

$$\Pr(A) = 1 - \Pr(\bar{A})$$

$$= 1 - \Pr(\bar{A}_1 \wedge \bar{A}_2 \wedge \bar{A}_3 \dots \wedge \bar{A}_n)$$

$$= 1 - \Pr(\bar{A}_1) \Pr(\bar{A}_2) \dots \Pr(\bar{A}_n)$$

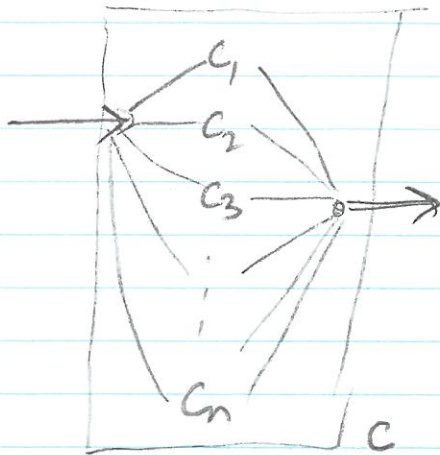
Mutual
Independence

$$= 1 - (1-p)^n$$

Thus for large values of n , $(1-p)^n \rightarrow 0$

$\Rightarrow \Pr(A) \rightarrow 1$, i.e. circuit will very likely fail.

Suppose C is configured as



Now C will fail only when all components fail.

$$\text{Thus } A = A_1 \wedge A_2 \wedge \dots \wedge A_n$$

$$\begin{aligned} \text{or } \Pr(A) &= \Pr(A_1 \wedge A_2 \wedge \dots \wedge A_n) \\ &= \Pr(A_1) \Pr(A_2) \dots \Pr(A_n) \\ &= p^n \end{aligned}$$

Mutual
Independence

Thus, when $n \rightarrow \infty$ $p^n \rightarrow 0$,

$\Rightarrow$ it is unlikely the circuit will fail
for large values of n .

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How to choose a random element from a linked list

Given:

A link list L



- We do not know $|L|$ -
- We are only allowed to make a single pass

Problem: Choose a node of L , uniformly at random.
 i.e., if $|L|=n$, then each node has $\text{pr} = 1/n$ of being chosen.

Algorithm ChooseRandom(L):

$u := \text{head}(L);$

$i := 1;$

While $u \neq \text{NIL}$ do

$r := \text{Random}(i);$ $\text{if } r = 1 \text{ then } x := u;$ $u := \text{succ}(u);$ $i := i + 1$	$\{ \text{Returns a random element with uniform probability from } \{1..i\} \}$
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Return (x).

Observe

- In the first iteration $x = \text{first element of } L$ with $p_r = 1$
- In the second iteration $x = \text{second element of } L$ with $p_r = 1/2$, and it does not change with $p_r = 1/2$
- In the third iteration $x = \text{third element of } L$ with $p_r = 1/3$; value of x does not change with $p_r = 2/3$.
- ⋮
- In the last iteration $x = \text{last element of } L$ with $p_r = 1/|L|$; and its value does not change with $p_r = |L|-1/|L|$

To Show: Out put of $\text{chooseRandom}(L)$ is a uniformly random node of List L . Let $|L| = n$, and let v be an arbitrary node of L . We will show, after the algorithm has terminated

$$x = v \text{ with } p_r = \frac{1}{|L|} = \frac{1}{n}.$$

Pf: Let v be the k^{th} node of L , $1 \leq k \leq n$.

Note that $x = v$ if and only if

- during the k^{th} iteration, the value of x is set to v .
- for all $i = k+1, k+2, \dots, n$; the value of x does not change in iteration i .

Define events

$$A = \{ \text{"After the algo has terminated } x = v" \}.$$

For $1 \leq i \leq n$, define

$$A_i = \{ \text{"the value of } x \text{ changes during the } i^{\text{th}} \text{ iteration"} \}$$

Then

$$A = A_k \cap \bar{A}_{k+1} \cap \bar{A}_{k+2} \cap \dots \cap \bar{A}_n$$

Since events $A_k, A_{k+1}, \dots, A_n$ are mutually independent.

$$\begin{aligned} \Pr(A) &= \Pr(A_k \cap \bar{A}_{k+1} \cap \bar{A}_{k+2} \cap \dots \cap \bar{A}_n) \\ &= \Pr(A_k) \Pr(\bar{A}_{k+1}) \Pr(\bar{A}_{k+2}) \dots \Pr(\bar{A}_n) \\ &= \frac{1}{k} \left(1 - \frac{1}{k+1}\right) \left(1 - \frac{1}{k+2}\right) \dots \left(1 - \frac{1}{n}\right) \\ &= \frac{1}{k} \cdot \frac{k}{k+1} \cdot \frac{k+1}{k+2} \cdot \dots \cdot \frac{n-1}{n} \\ &= \frac{1}{n} \quad \blacksquare \end{aligned}$$